

**On the Random Walk with Zero Drifts  
in the First Quadrant of  $\mathbb{R}_2$**

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**ABSTRACT**

The random walk on the lattice in the positive quadrant is studied for the case that the one-step displacement vector has zero drift, finite second moments and support contained in  $(-1, 0, 1, 2, \dots) \times (-1, 0, 1, 2, \dots)$ . It is well known that the first entrance time out from a point in the interior of the lattice into the union of the coordinate axes is finite with probability one. In the present study it is shown that its first moment is finite if and only if the covariance of the  $x$ - and  $y$ -component of the one-step displacement vector is negative. For this case explicit expressions are given for the first moment of the entrance time and for the first and second moments of the hitting point of the axes, in terms of the second moment characteristics of the one-step displacement vector. The results are deduced from the hitting point identity for the random walk.

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## 1. INTRODUCTION

The random walk on the lattice with nonnegative, integer valued coordinates in  $\mathbb{R}_2$  is studied for the case that the  $x$ - and  $y$ -component of the one-step displacement vector have zero drift, and finite second moments, the support of the vector being contained in  $\{-1, 0, 1, \dots\} \times \{-1, 0, 1, \dots\}$ . It is well known that the first entrance time out from a point in the interior of the lattice into the set formed by the coordinate axes is finite with probability one. In the present study necessary and sufficient conditions for the first moment of this entrance time to be finite are derived. In the case that it is finite an explicit expression for this moment is obtained, similarly for the first and second moments of the hitting point at the coordinate axes. The starting point of the derivations is the so called hitting point identity, it has been derived in a somewhat more general context in Cohen, [1], [2], and it represents a relation for the bivariate generating function of the joint distribution at the first entrance time and the hitting point on the set of zero tuples of the so called kernel of the two-dimensional random walk, see formules (1.12) and (1.13) below.

The hitting point identity is actually the key relation for the analysis of two-dimensional random walks on the first quadrant. In general its decomposition can be reduced to the solution of a Riemann Boundary Value problem, cf. Cohen [3], this leads to a rather intricate analysis. However, in the present case with zero drifts it is possible to deduce directly from the hitting point identity explicit results. For the derivations here we need some rather basic asymptotical results concerning the zero tuples. These asymptotical results have been derived in the appendix of the present paper; they play also an important role in the study of the ergodicity conditions of the random walk with reflecting boundaries, see Cohen [4].

The main result of the present study is formulated in theorem 1.1, next to this it shows the importance of the hitting point identity.

An elegant derivation of the statement (1.10)iii of theorem 1.1 has been presented by Klein Haneveld and Pittenger [6] and a complete proof of the theorem by Durrett [5]. These derivations, which are fairly simple, have been obtained by using martingale theorems. However, the approach via martingales fails for the analogous higher dimensional random walks. Here the approach via the hitting point identity leads also to results as it will be shown in a forthcoming study.

The random walk  $z_n \equiv (x_n, y_n)$ ,  $n = 0, 1, 2, \dots$ , with state space  $S$ , the lattice points with integer valued, nonnegative, coordinates in  $\mathbb{R}_2$ , i.e.

$$S = \{0, 1, 2, \dots\} \times \{0, 1, 2, \dots\}, \quad (1.1)$$

is defined by: for  $n = 0, 1, 2, \dots$ ,

i. for  $x_n > 0$ ,  $y_n > 0$ :

$$\begin{aligned} x_{n+1} &= x_n - 1 + \xi_n, \\ y_{n+1} &= y_n - 1 + \eta_n; \end{aligned} \quad (1.2)$$

ii. for  $x_n = 0$  or  $y_n = 0$ ,

$$x_{n+1} = x_n,$$

$$y_{n+1} = y_n;$$

iii.  $x_0 = x_0, y_0 = y_0$ ;

with

i. the starting point

$$z_0 \equiv (x_0, y_0) \in S, \tag{1.3}$$

ii.  $(\xi_n, \eta_n), n = 0, 1, 2, \dots$ , a sequence of i.i.d. vectors each with state space  $S$  and

$$E\{\xi_n\} = E\{\eta_n\} = 1, \quad \alpha_1 := E\{(\xi_n - 1)^2\} < \infty, \quad \alpha_2 := E\{(\eta_n - 1)^2\} < \infty,$$

$$\alpha_{12} := E\{(\xi_n - 1)(\eta_n - 1)\}.$$

Let  $(\xi, \eta)$  be a stochastic vector with the same state space and bivariate distribution as  $(\xi_n, \eta_n)$ , i.e.

$$(\xi, \eta) \sim (\xi_n, \eta_n).$$

The bivariate generating function of the distribution of  $(\xi, \eta)$  is defined by

$$\phi(p_1, p_2) := E\{p_1^\xi p_2^\eta\}, \quad |p_1| \leq 1, |p_2| \leq 1. \tag{1.4}$$

We introduce the

*Assumption.*

i.  $|\phi(p_1, p_2)| = 1$  for  $|p_1| = 1, |p_2| = 1 \Rightarrow p_2 = 1, p_1 = 1$ ;  $\tag{1.5}$

ii. for every  $(i, j) \in S$  the coefficient of  $p_1^i p_2^j$  in the series expansion of  $[\phi(p_1, p_2)/p_1 p_2]^n$  with  $|p_1| = 1, |p_2| = 1, n$  being a positive integer, is positive for  $n$  sufficiently large;

iii.  $\phi(0, 1) > 0, \phi(1, 0) > 0$ .

REMARK 1.1. (1.5)i implies that the random walk  $z_n$  is aperiodic, cf. Spitzer [7]; whereas (1.5)ii implies that its state space  $S$  is irreducible; (1.5)iii is an obvious assumption, otherwise  $x_n$  and/or  $y_n$  cannot decrease.  $\square$

Put

$$B := \{0\} \times \{0, 1, 2, \dots\} \cup \{0, 1, 2, \dots\} \times \{0\}, \tag{1.6}$$

so  $B$  is the boundary of  $S$ .

Define

$$\begin{aligned}
\mathbf{m}(z_0) &:= \inf_{n=0,1,\dots} \{n : \mathbf{z}_n \in B | z_0 \in S\}, & (1.7) \\
&:= \infty \text{ if } \mathbf{z}_n \notin B \text{ for all } n=0,1,2,\dots; \\
\mathbf{k}(z_0) \equiv (\mathbf{k}_1(z_0), \mathbf{k}_2(z_0)) &:= \mathbf{z}_{\mathbf{m}(z_0)} \equiv (\mathbf{x}_{\mathbf{m}(z_0)}, \mathbf{y}_{\mathbf{m}(z_0)}) \text{ if } \mathbf{m}(z_0) < \infty, \\
&:= (\infty, \infty) \text{ ,, } = \infty.
\end{aligned}$$

So  $\mathbf{m}(z_0)$  is the *first entrance time* of the  $\mathbf{z}_n$ -processes into the boundary  $B$  when starting at  $z_0$ , and  $\mathbf{k}(z_0)$  is then the *hitting point* of  $B$ , note that

$$\mathbf{k}_1(z_0)\mathbf{k}_2(z_0) = 0 \quad \text{if } \mathbf{m}(z_0) < \infty. \quad (1.9)$$

The main result of the present study is the following

**THEOREM 1.1.** For  $z_0 \in S \setminus B$ :

- i.  $\Pr\{\mathbf{m}(z_0) < \infty\} = 1$ ;
  - ii.  $E\{\mathbf{k}_1(z_0)\} = x_0, E\{\mathbf{k}_2(z_0)\} = y_0$ ; (1.10)
  - iii.  $E\{(\xi-1)(\eta-1)\} < 0 \Leftrightarrow E\{\mathbf{m}(z_0)\} < \infty, E\{\mathbf{k}_1^2(z_0)\} < \infty, E\{\mathbf{k}_2^2(z_0)\} < \infty$ ;
- and for  $E\{(\xi-1)(\eta-1)\} < 0$ ,

- i.  $E\{\mathbf{m}(z_0)\} = \frac{x_0 y_0}{-E\{(\xi-1)(\eta-1)\}}$ ;
- ii.  $E\{(\mathbf{k}_1(z_0) - x_0)^2\} = E\{(\xi-1)^2\} E\{\mathbf{m}(z_0)\},$  (1.11)  
 $E\{(\mathbf{k}_2(z_0) - y_0)^2\} = E\{(\eta-1)^2\} E\{\mathbf{m}(z_0)\}.$

Theorem 1.1 is quite remarkable since it formulates explicitly in terms of the second moments of the one-step displacement vector  $(\xi-1, \eta-1)$  the necessary and sufficient condition for the first moment of the entrance time to be finite and further the expression for this moment as well as the expressions for the first and second moments of the hitting point. Actually (1.10)i is a wellknown result for two-dimensional random walks for which (1.3)i applies, cf. Spitzer [7].

The proof of theorem 1.1 is based on the hitting point identity for the  $\mathbf{z}_n$ -process, see below, and it is presented in Section 2. The proof needs several asymptotic results, these are derived in the Appendix, see Section 3.

**REMARK 1.2.** Mostly we shall delete  $z_0$  in the symbols defined in (1.7) and (1.8), i.e. we shall write for  $z_0 \in S \setminus B$ ,

$$\mathbf{m} \equiv \mathbf{m}(z_0), \quad \mathbf{k}_1 \equiv \mathbf{k}_1(z_0), \quad \mathbf{k}_2 \equiv \mathbf{k}_2(z_0);$$

further  $(A)$  shall denote the indicator function of the event  $A$ , i.e.

$$(A) = 1 \text{ if } A \text{ occurs,} \\ = 0 \text{ ,, } \bar{A} \text{ ,, .} \quad \square$$

The *hitting point identity* for the  $z_n$ -process defined above reads: for  $|r| \leq 1$ ,

$$\hat{p}_1^{x_0} \hat{p}_2^{y_0} = E\{r^m \hat{p}_1^{k_1} (\mathbf{k}_1 > 0)\} + E\{r^m (\mathbf{k}_1 = \mathbf{k}_2 = 0)\} + E\{r^m \hat{p}_2^{k_2} (\mathbf{k}_2 > 0)\}, \quad (1.12)$$

with for  $|r| \leq 1, r \neq 1$ ,

$(\hat{p}_1, \hat{p}_2)$  any zero tuple of the *kernel*

$$Z(r, p_1, p_2) := p_1 p_2 - r \phi(p_1, p_2), \quad |p_1| \leq 1, |p_2| \leq 1, \quad (1.13)$$

and for  $r = 1$ ,

$(\hat{p}_1, \hat{p}_2)$  any zero tuple which is the limit for  $r \rightarrow 1$  of a zero tuple  $(\hat{p}_1, \hat{p}_2)$  of  $Z(r, p_1, p_2)$  with  $|r| \leq 1, r \neq 1$ .

REMARK 1.3.  $(\hat{p}_1, \hat{p}_2)$  is said to be a zero tuple of  $Z(r, p_1, p_2)$  if

$$Z(r, \hat{p}_1, \hat{p}_2) = 0.$$

REMARK 1.4. The H.P.I. (1.10) is a special case of a more general identity discussed in Cohen [1], viz. take in the notation of [1],  $N = 2, M = 0$  and  $T$  empty.

## 2. PROOF OF THEOREM 1.1

From lemma 3.1 it follows that  $(\hat{p}_1, \hat{p}_2)$  with

$$\hat{p}_2 = p_2, |p_2| = 1, \hat{p}_1 = P_1(r, p_2), |r| \leq 1, r \neq 1, \quad (2.1)$$

is a zero tuple of  $Z(r, p_1, p_2), |p_1| \leq 1, |p_2| \leq 1$ , and that for  $r \rightarrow 1$  it is also a zero tuple. By symmetry the zero tuple

$$(p_1, P_2(r, p_1)), \quad |p_2| = 1,$$

is constructed.

Hence from (1.12) we have for  $z_0 \in S \setminus B, |r| \leq 1, |p_2| = 1$ ,

$$P_1^{x_0}(r, p_2) p_2^{y_0} = E\{r^m P_1^{k_1}(r, p_2) (\mathbf{k}_1 \geq 0, \mathbf{k}_2 = 0)\} + E\{r^m p_2^{k_2} (\mathbf{k}_2 > 0)\}. \quad (2.2)$$

Since  $P_1(r, 1) \rightarrow 1$  for  $r \rightarrow 1$  it follows from (2.2) by taking  $p_2 = 1$  and letting  $r \rightarrow 1$  that

$$1 = \Pr\{\mathbf{m} < \infty\},$$

i.e. (1.10) i has been proved. For  $|r| \leq 1, r \neq 1, |p_2| = 1$  we have  $|P_1(r, p_2)| < 1$  so

$$|E\{r^m \mathbf{k}_1 P_1^{k_1-1}(r, p_2)\}| < \infty;$$

since  $P_1(r, p_2)$  possesses a derivative with respect to  $p_2$ , see lemma 3.1ii, it follows that the last term in (2.2) possesses also a first derivative.

Differentiate (2.2) with respect to  $p_2$  and then take  $p_2 = 1$ , it then follows for  $|r| \leq 1, r \neq 1$ ,

$$\frac{\partial P_1(r, p_2)}{\partial p_2} \Big|_{p_2=1} [x_0 P_1^{x_0-1}(r, 1) - E\{r^m \mathbf{k}_1 P_1^{\mathbf{k}_1-1}(r, 1)\} + y_0 P_1^{y_0}(r, 1) - E\{r^m \mathbf{k}_2\}] = \tag{2.3}$$

So by letting  $r \rightarrow 1$  we obtain from (1.3), (2.3) and (3.9),

$$-\frac{\alpha_{12}}{\alpha_1} [x_0 - E\{\mathbf{k}_1\}] + y_0 - E\{\mathbf{k}_2\} = 0, \tag{2.4}$$

$$x_0 - E\{\mathbf{k}_1\} - \frac{\alpha_{12}}{\alpha_2} [y_0 - E\{\mathbf{k}_2\}] = 0;$$

where the second relation in (2.4) follows by symmetry, cf. remark 2.1, from the first one. Because of (3.3) the main determinant of the set of equations is nonzero, so this system has only the zero solution, if  $E\{\mathbf{k}_1\} < \infty, E\{\mathbf{k}_2\} < \infty$ , and then (1.10)ii follows; for the ultimate proof see below, directly after (2.13), and note that that proof uses only (2.2) and (3.6).

By taking  $p_2 = 1$  in (2.2) and differentiating with respect to  $r$  we obtain for  $|r| \leq 1, r \neq 1$ ,

$$\frac{dP_1(r, 1)}{dr} [x_0 P_1^{x_0-1}(r, 1) - E\{r^m \mathbf{k}_1 P_1^{\mathbf{k}_1-1}(r, 1)\}] = \tag{2.5}$$

$$E\{\mathbf{m}r^{\mathbf{m}-1} P_1^{\mathbf{k}_1}(r, 1) (\mathbf{k}_1 \geq 0, \mathbf{k}_2 = 0)\} + E\{\mathbf{m}r^{\mathbf{m}-1} (\mathbf{k}_2 > 0)\}.$$

From (2.3) and (2.5) it follows for  $|r| \leq 1, r \neq 1$ ,

$$\frac{\partial P_1(r, p_2)}{\partial p_2} \Big|_{p_2=1} [E\{\mathbf{m}r^{\mathbf{m}-1} \{P_1^{\mathbf{k}_1}(r, 1) (\mathbf{k}_1 \geq 0, \mathbf{k}_2 = 0) + (\mathbf{k}_2 > 0)\}\}] = \tag{2.6}$$

$$-\frac{dP_1(r, 1)}{dr} [y_0 P_1^{y_0}(r, 1) - E\{r^m \mathbf{k}_2\}].$$

We consider this relation for  $|r| \leq 1, r \sim 1$ , it then follows from (2.6), (3.5)ii, (3.9) and (1.10)ii that

$$\{-\alpha_{12} + o(1)\} E\{\mathbf{m}r^{\mathbf{m}-1} [P_1^{\mathbf{k}_1}(r, 1) (\mathbf{k}_1 \geq 0, \mathbf{k}_2 = 0) + (\mathbf{k}_2 > 0)]\} = \tag{2.7}$$

$$\left\{-2 + \frac{1}{2} \sqrt{2} \frac{\alpha_1^{\frac{1}{2}}}{\sqrt{1-r}} \left(1 - \frac{o(\sqrt{1-r})}{\sqrt{1-r}}\right)\right\} [y_0 (1 - P_1^{x_0}(r, 1)) - E\{\mathbf{k}_2 (1 - r^{\mathbf{m}})\}]$$

$$+ o(\sqrt{1-r}) = -2 \sqrt{2} \sqrt{1-r} x_0 y_0 \alpha_1^{\frac{1}{2}} + 2 E\{\mathbf{k}_2 (1 - r^{\mathbf{m}})\} +$$

$$[x_0 y_0 - \frac{1}{2} \sqrt{2} \alpha_1^{\frac{1}{2}} E\{\mathbf{k}_2 \frac{1-r^{\mathbf{m}}}{\sqrt{1-r}}\}] \left(1 - \frac{o(\sqrt{1-r})}{\sqrt{1-r}}\right).$$

Suppose

$$\alpha_{12} < 0. \tag{2.8}$$

then for  $r > 0$  and  $r \sim 1$  the lefthand side of (2.7) is positive, possibly  $+\infty$  for  $r \rightarrow 1$ , and since

$$E\left\{k_2 \frac{1-r^m}{\sqrt{1-r}}\right\} > 0 \text{ for } 0 < r < 1,$$

it follows from (2.7) that

$$x_0 y_0 \sqrt{2} \alpha_1^{-1/2} > \limsup_{r \rightarrow 1} E\left\{k_2 \frac{1-r^m}{\sqrt{1-r}}\right\} \geq 0.$$

Hence the lefthand side of (2.7) has for  $r \rightarrow 1$  a finite limit which since  $P_1(r, 1) \rightarrow 1$  is equal to

$$-\alpha_{12} E\{\mathbf{m}\},$$

and hence  $E\{\mathbf{m}\}$  should be finite, so we have shown that, cf. (2.8),

$$\alpha_{12} < 0 \Rightarrow E\{\mathbf{m}\} < \infty, \tag{2.9}$$

and moreover that the following limit exists and

$$x_0 y_0 \sqrt{2} \alpha_1^{-1/2} > \lim_{r \rightarrow 1} E\left\{k_2 \frac{1-r^m}{\sqrt{1-r}}\right\} \geq 0. \tag{2.10}$$

Next we take in (2.2)  $r=1$  and differentiate with respect to  $p_2$ , cf. lemma 3.1 iv, this yields for  $|p_2|=1, p_2 \neq 1$ ,

$$E\{k_1 P_1^{k_1-1}(1, p_2)\} \frac{dP_1(1, p_2)}{dp_2} + E\{k_2 p_2^{k_2-1}\} = \tag{2.11}$$

$$x_0 P_1^{x_0-1}(1, p_2) p_2^{y_0} \frac{dP_1(1, p_2)}{dp_2} + y_0 P_1^{x_0}(1, p_2) p_2^{y_0-1},$$

and

$$\begin{aligned} & E\{k_1(k_1-1)P_1^{k_1-2}(1, p_2)\} \left\{ \frac{dP_1(1, p_2)}{dp_2} \right\}^2 + \\ & E\{k_1 P_1^{k_1-1}(1, p_2)\} \frac{d^2 P_1(1, p_2)}{dp_2^2} + E\{k_2(k_2-1)p_2^{k_2-2}\} = \\ & x_0(x_0-1)P_1^{x_0-2}(1, p_2) p_2^{y_0} \left[ \frac{dP_1(1, p_2)}{dp_2} \right]^2 + x_0 P_1^{x_0-1}(1, p_2) p_2^{y_0} \frac{d^2 P_2(1, p_2)}{dp_2^2} + \\ & 2x_0 y_0 P_1^{x_0-1}(1, p_2) p_2^{y_0-1} \frac{dP_1(1, p_2)}{dp_2} + y_0(y_0-1)P_1^{x_0}(1, p_2) p_2^{y_0-2}, \end{aligned} \tag{2.12}$$

as in the derivation of (2.5) it is readily shown that  $E\{p_2^{k_2}(k_2 > 0)\}$  is

twice differentiable for  $|p_2|=1, p_2 \neq 1$ . Letting  $p_2 \rightarrow 1$  in (2.11) yields since  $P_1(1, p_2) \rightarrow 1$ ,

$$\{E\{\mathbf{k}_1\} - x_0\} \frac{dP_1(1, p_2)}{dp_2} \Big|_{p_2=1} = y_0 - E\{\mathbf{k}_2\}. \tag{2.13}$$

From (3.3) and (3.6) it is seen that  $\frac{dP_1(1, p_2)}{dp_2} \Big|_{p_2=1}$  is complex and hence we obtain (1.10)ii.

Suppose

$$E\{\mathbf{k}_1^2\} < \infty, \quad E\{\mathbf{k}_2^2\} < \infty. \tag{2.14}$$

It then follows from (2.12) that for  $p_2 \rightarrow 1$ ,

$$[E\{\mathbf{k}_1(\mathbf{k}_1 - 1)\} - x_0(x_0 - 1)] \left\{ \frac{dP_1(1, p_2)}{dp_2} \Big|_{p_2=1} \right\}^2 + \tag{2.15}$$

$$E\{\mathbf{k}_2(\mathbf{k}_2 - 1)\} - y_0(y_0 - 1) - 2x_0y_0 \frac{dP_1(1, p_2)}{dp_2} \Big|_{p_2=1} =$$

$$- \lim_{p_2 \rightarrow 1} \left\{ [E\{\mathbf{k}_1 P_1^{\mathbf{k}_1 - 1}(1, p_2)\} - x_0 P_1^{x_0 - 1}(1, p_2) p_2^{y_0}] \frac{d^2 P_1(1, p_2)}{dp_2^2} \right\},$$

where the limit in the righthand side should exist since (2.14) and (3.6) show that the lefthand side is finite.

By using (1.10)ii, i.e.  $E\{\mathbf{k}_1\} = x_0$ , it follows that for  $p_2 \rightarrow 1$ ,

$$E\{\mathbf{k}_1 P_1^{\mathbf{k}_1 - 1}(1, p_2)\} - x_0 P_1^{x_0 - 1}(1, p_2) p_2^{y_0} \rightarrow 0, \tag{2.16}$$

and so since  $\frac{dP_1(1, p_2)}{dp_2} \Big|_{p_2=1}$  is finite it follows from (2.14) that the left hand side of (2.16) behaves as  $1 - p_2$  for  $p_2 \sim 1, |p_2|=1$ . Consequently, it follows from lemma 3.3ii that the righthand side of (2.15) is equal to zero; by using (3.6)i, or better (3.8), a simple calculation shows that (2.15) with its righthand side being zero may be rewritten as

$$2[E\{(\mathbf{k}_1 - x_0)^2\} \frac{\alpha_{12}}{\alpha_1} + x_0 y_0] \frac{dP_1(1, p_2)}{dp_2} \Big|_{p_2=1} = E\{(\mathbf{k}_2 - y_0)^2\} - \frac{\alpha_2}{\alpha_1} E\{(\mathbf{k}_1 - x_0)^2\}. \tag{2.17}$$

Since  $\frac{dP_1(1, p_2)}{dp_2} \Big|_{p_2=1}$  is complex it follows from (2.17) that

$$E\{(\mathbf{k}_1 - x_0)^2\} = x_0 y_0 \frac{\alpha_1}{-\alpha_{12}}, \quad E\{(\mathbf{k}_2 - y_0)^2\} = x_0 y_0 \frac{\alpha_2}{-\alpha_{12}} \tag{2.18}$$

The lefthand sides in (2.18) are positive and consequently it has been shown that, cf. (2.14),

$$E\{\mathbf{k}_1^2\} < \infty, \quad E\{\mathbf{k}_2^2\} < \infty \Rightarrow \alpha_{12} < 0. \tag{2.19}$$



To complete the proof of (1.10)iii it is seen from (2.9) and (2.19) that we have still to show that

$$E\{\mathbf{m}\} < \infty \Rightarrow E\{\mathbf{k}_1^2\} < \infty, \quad E\{\mathbf{k}_2^2\} < \infty. \tag{2.20}$$

To prove (2.20) we rewrite (2.5) for  $|r| \leq 1, r \neq 1$  as

$$E\{\mathbf{m}r^m P_1^{\mathbf{k}_1}(r, 1)(\mathbf{k}_1 \geq 0, \mathbf{k}_2 = 0)\} + E\{\mathbf{m}r^{m-1}(\mathbf{k}_2 > 0)\} = \tag{2.21}$$

$$\frac{dP_1(r, 1)}{dr} (1 - P_1(r, 1)) \left[ -x_0 \frac{1 - P_1^{x_0-1}(r, 1)}{1 - P_1(r, 1)} + E\{\mathbf{k}_1 r^m \frac{1 - P_1^{\mathbf{k}_1-1}(r, 1)}{1 - P_1(r, 1)}\} \right] +$$

$$\frac{dP_1(r, 1)}{dr} \sqrt{1-r} E\left\{\mathbf{k}_1 \frac{1-r^m}{\sqrt{1-r}}\right\}.$$

Hence by using the asymptotic relations (3.5) we have from (2.21) for  $r \sim 1$ ,

$$E\{\mathbf{m}r^m P_1^{\mathbf{k}_1}(r, 1)(\mathbf{k}_1 \geq 0, \mathbf{k}_2 = 0)\} + E\{\mathbf{m}r^{m-1}(\mathbf{k}_2 > 0)\} = \tag{2.22}$$

$$E\left\{\mathbf{k}_1 \frac{1-r^m}{\sqrt{1-r}}\right\} \left\{ -2\alpha_1^{-1} \sqrt{1-r} + \frac{1}{2} \sqrt{2} \alpha^{-1/2} \left(1 - \frac{o(\sqrt{1-r})}{\sqrt{1-r}}\right) + o(\sqrt{1-r}) \right\} +$$

$$\left\{ -2\sqrt{2} \alpha_1^{-1/2} \sqrt{1-r} + \alpha_1^{-1} \left(1 - \frac{o(\sqrt{1-r})}{\sqrt{1-r}}\right) \right\}$$

$$- x_0 \frac{1 - P_1^{x_0-1}(r, 1)}{1 - P_1(r, 1)} + E\left\{\mathbf{k}_1 r^m \frac{1 - P_1^{\mathbf{k}_1-1}(r, 1)}{1 - P_1(r, 1)}\right\}.$$

We next observe that (2.7) implies that

$$E\{\mathbf{m}r^{m-1} P_1^{\mathbf{k}_1}(r, 1)(\mathbf{k}_1 \geq 0, \mathbf{k}_2 = 0)\} + E\{\mathbf{m}r^{m-1}(\mathbf{k}_2 > 0)\} > 0, \quad 0 < r < 1.$$

and

$$E\left\{\mathbf{k}_2 \frac{1-r^m}{\sqrt{1-r}}\right\} > 0, \quad 0 < r < 1,$$

both have a limit for  $r \rightarrow 1$  and these limits are both finite or both  $+\infty$ . So by symmetry, cf. remark 2.1, the same holds for

$$E\{\mathbf{m}r^{m-1} P_2^{\mathbf{k}_2}(r, 1)(\mathbf{k}_2 \geq 0, \mathbf{k}_1 = 0)\} + E\{\mathbf{m}r^m(\mathbf{k}_1 > 0)\}$$

and

$$E\left\{\mathbf{k}_1 \frac{1-r^m}{\sqrt{1-r}}\right\}.$$

Suppose

$$E\{\mathbf{m}\} < \infty, \tag{2.23}$$

then by the observation just made it follows from (2.22) that the following limit should exist and be finite:

$$|\lim_{r \rightarrow 1} \{ E\{k_1 r^m \frac{1 - P_1^{k_1 - 1}(r, 1)}{1 - P_1(r, 1)}\} - x_0 \frac{1 - P_1^{x_0 - 1}(r, 1)}{1 - P_1(r, 1)} \}| < \epsilon \tag{2.24}$$

Since  $k_1(k_1 - 1) \geq 0$  with probability one it follows that for  $0 < r < 1$  and  $r \rightarrow 1$ ,

$$E\{k_1 r^m \frac{1 - P_1^{k_1 - 1}(r, 1)}{1 - P_1(r, 1)}\} \leq E\{k_1 \frac{1 - P_1^{k_1 - 1}(r, 1)}{1 - P_1(r, 1)}\} \rightarrow E\{k_1(k_1 - 1)\},$$

$$x_0 \frac{1 - P_1^{x_0 - 1}(r, 1)}{1 - P_1(r, 1)} \rightarrow x_0(x_0 - 1), \tag{2.25}$$

and consequently (2.23) implies that  $E\{k_1^2\}$  should be finite, and by symmetry also  $E\{k_2^2\} < \infty$ . Hence (2.20) has been proved, and so is (1.10)iii.

To prove (1.11) let  $r \rightarrow 1$  in (2.7) and (2.22) with  $\alpha_{12} < 0$  then it follows

$$-\alpha_{12} E\{m\} = x_0 y_0 - \frac{1}{2} \sqrt{2} \alpha_1^{1/2} \lim_{r \rightarrow 1} E\{k_2 \frac{1 - r^m}{\sqrt{1 - r}}\}, \tag{2.26}$$

$$= x_0 y_0 - \frac{1}{2} \sqrt{2} \alpha_2^{1/2} \lim_{r \rightarrow 1} E\{k_1 \frac{1 - r^m}{\sqrt{1 - r}}\},$$

$$E\{m\} = \frac{1}{2} \sqrt{2} \alpha_1^{-1/2} \lim_{r \rightarrow 1} E\{k_1 \frac{1 - r^m}{\sqrt{1 - r}}\} + \alpha_1^{-1} E\{(k_1 - x_0)^2\}, \tag{2.27}$$

with the second relation in (2.26) based on symmetry. Elimination of the limit from (2.26) and (2.27) yields by using (2.18),

$$E\{m\} = \frac{x_0 y_0}{-\alpha_{12}}, \tag{2.28}$$

$$\lim_{r \rightarrow 1} E\{k_1 \frac{1 - r^m}{\sqrt{1 - r}}\} = \lim_{r \rightarrow 1} E\{k_2 \frac{1 - r^m}{\sqrt{1 - r}}\} = 0,$$

and hence (1.11) follows from (2.27) and (2.28). □

### 3. APPENDIX

For the proof of theorem 1.1 we need a number of properties of a special class of zero tuples of the kernel, i.e. for  $|r| \leq 1$ , of

$$Z(r, p_1, p_2) := p_1 p_2 - r E\{p_1^\xi p_2^\eta\}, \quad |p_1| \leq 1, |p_2| \leq 1, \tag{3.1}$$

with, cf. (1.3)ii,

$$E\{\xi\} = 1, \quad E\{\eta\} = 1, \quad E\{\xi^2\} < \infty, \quad E\{\eta^2\} < \infty.$$

Put

$$\alpha_1 := E\{(\xi - 1)^2\}, \quad \alpha_{12} := E\{(\xi - 1)(\eta - 1)\}, \quad \alpha_2 := E\{(\eta - 1)^2\},$$

so that, cf. (1.5),

$$\alpha_1 \alpha_2 > \alpha_1^2. \tag{3.3}$$

LEMMA 3.1. *i. The kernel  $Z(r, p_1, p_2)$  has for  $|r| \leq 1, r \neq 1, |p_2| = 1$ , a unique zero, say,  $P_1(r, p_2)$  in  $|p_1| \leq 1$ ;*

*ii.  $P_1(1, p_2) := \lim_{r \rightarrow 1} P_1(r, p_2), |p_2| = 1, p_2 \neq 1$  is a unique zero of  $Z(1, p_1, p_2)$  in  $|p_1| \leq 1$ ,*

$$|P_1(1, p_2)| < 1 \text{ for } |p_2| = 1, |p_2| \neq 1,$$

$$P_1(1, 1) := \lim_{p_2 \rightarrow 1} P_1(1, p_2) = 1,$$

*and all these zeros have multiplicity one except  $P_1(1, 1)$  which has multiplicity two.*

*iii.  $P_1(r, p_2), |r| \leq 1, r \neq 1$  is a twice differentiable function of  $p_2$  on  $|p_2| = 1$ , the same applies for  $P_1(1, p_2)$  on  $|p_2| = 1, p_2 \neq 1$ .*

*iv.  $P_1(r, p_2), |p_2| = 1$  is a regular function of  $r$  for  $|r| \leq 1, r \neq 1$ .*

PROOF. For fixed  $p_2$  with  $|p_2| = 1$  it is easily seen that  $E\{p_1^\xi p_2^{\eta-1}\}$  is a regular function of  $p_1$  for  $|p_1| < 1$ , and continuous for  $|p_1| \leq 1$ . From (1.5)i it is seen for  $|r| \leq 1, r \neq 1, |p_2| = 1$  that on  $|p_1| = 1$ ,

$$|p_1| = 1 > |r| |E\{p_1^\xi p_2^{\eta-1}\}|,$$

so that by applying Rouché's theorem the first statement follows, and  $P_1(r, p_2), |r| \leq 1, r \neq 1, |p_2| = 1$  is a single zero and is nonzero for  $r \neq 0$ , cf. (1.5)iii. Obviously,  $P_1(r, p_2), |r| < 1, |p_2| = 1$  is a continuous function of  $r$  since  $Z(r, p_1, p_2)$  is such a function, so  $Z(1, p_1, p_2), |p_2| = 1$  can have at most one zero in  $|p_1| < 1$ , and since for  $|p_1| = 1, p_1 \neq 1, |p_2| = 1$ ,

$$|p_1| > |E\{p_1^\xi p_2^{\eta-1}\}|,$$

it follows that  $P_1(1, p_2)$  is the only zero of  $Z(1, p_1, p_2), |p_2| = 1, p_2 \neq 1$  in  $|p_1| \leq 1$ , it has multiplicity one and is nonzero, cf. (1.5)iii. Obviously  $p_1 = 1$  is a zero of  $Z(1, p_1, 1)$ , so that the continuity of  $Z(1, p_1, p_2)$  in  $p_1, p_2$  on  $|p_1| \leq 1, |p_2| = 1$  implies that  $P_1(1, p_2) \rightarrow 1$  for  $p_2 \rightarrow 1$ . Because  $E\{\xi\} = 1$  and  $P_1(r, 1)$  is positive for  $0 < r \leq 1$  it is readily verified that  $P_1(1, 1)$  is a zero with multiplicity two, note that

$$\lim_{r \rightarrow 1} P_1(r, 1) = \lim_{p_2 \rightarrow 1} P_1(1, p_2).$$

The third statements follows directly from the fact that  $Z(r, p_1, p_2)$  possess a derivative with respect to  $p_1$  for  $|p_1| \leq 1$ , and similarly with respect to  $p_2$  for  $|p_2| \leq 1$ , and the fact that  $P_1(r, p_2), |p_2| = 1$  and

$P_1(1, p_2)$ ,  $|p_2|=1$ ,  $p_2 \neq 1$  are zeros with multiplicity one. The proof of the  $i$ th statement is similar.  $\square$

REMARK 3.1. Denote by  $m_1$  for the component random walk  $x_n$ ,  $n=0, 1, \dots$ , cf. (1.2), with  $x_0=1$  the first entrance time into the zero state. Simple arguments from the theory of one dimensional random walks lead directly to

$$P_1(r, p_2) = E\{r^{m_1} p_2^{\sum_{h=1}^{m_1} (\eta_h - 1)}\}. \tag{3.4}$$

LEMMA 3.2. For  $|r| \leq 1$ ,  $r \sim 1$ ,

i.  $P_1(r, 1) = 1 - \sqrt{2} \alpha_1^{-1/2} \sqrt{1-r} + o(\sqrt{1-r})$ , (3.5)

ii.  $\frac{dP_1(r, 1)}{dr} = -2\alpha_1^{-1} + \frac{1}{2} \sqrt{2} \alpha_1^{-1/2} \frac{1}{\sqrt{1-r}} (1 - \frac{o(\sqrt{1-r})}{\sqrt{1-r}}) + O(\sqrt{1-r})$ .

PROOF. For  $p_2=1$  it follows from (3.1) and lemma 3.1 for  $|r| \leq 1$ ,  $r \neq 1$ , since  $P_1(1, 1)=1$  and  $\alpha_1 = E\{(\xi-1)^2\} < \infty$  that with  $r \sim 1$ ,

$$1 - (1 - P_1(r, 1)) = r E\{[1 - (1 - P_1(r, 1))]^\xi\} =$$

$$rE\{1\} - rE\{\xi\}(1 - P_1(r, 1)) + \frac{r}{2} E\{\xi(\xi - 1)\}(1 - P_1(r, 1))^2 + o((1 - P_1(r, 1))^2),$$

or with  $E\{\xi\} = 1$ ,

$$\frac{1}{2} r \alpha_1 (1 - P_1(r, 1))^2 + (1 - r)(1 - P_1(r, 1)) - (1 - r) + o((1 - P_1(r, 1))^2) = 0,$$

i.e.

$$\frac{1}{2} r \alpha_1 \left\{ \frac{1 - P_1(r, 1)}{\sqrt{1-r}} \right\}^2 + \sqrt{1-r} \frac{1 - P_1(r, 1)}{\sqrt{1-r}} - 1 + o\left(\left(\frac{1 - P_1(r, 1)}{\sqrt{1-r}}\right)^2\right) = 0.$$

Hence it follows that

$$\frac{1 - P_1(r, 1)}{\sqrt{1-r}} = O(1) \text{ for } r \sim 1,$$

and so, note that  $1 - P_1(r, 1) > 0$  for  $0 < r < 1$ ,

$$\frac{1 - P_1(r, 1)}{\sqrt{1-r}} = \frac{1}{r \alpha_1} \{-\sqrt{1-r} + \sqrt{1-r + 2r \alpha_1}\} + O(\sqrt{1-r}),$$

and (3.5)i follows.

Obvious  $P_1(r, 1)$  has a derivative for  $|r| \leq 1$ ,  $r \neq 1$ , and by using the fact that  $P_1(r, 1)$  is a zero of  $Z(r, p_1, 1)$  it follows readily for  $|r| \leq 1$ ,  $r \neq 1$ ,

$$\frac{dP_1(r, 1)}{dr} E\{(1 - \xi)P_1^\xi(r, 1)\} = \frac{1}{r} E\{P_1^{\xi+1}(r, 1)\},$$

or

$$\frac{dP_1(r, 1)}{dr} E\{(1-\xi)[1-(1-P_1(r, 1))]^\xi\} = \frac{1}{r} E\{[1-(1-P_1(r, 1))]^{\xi+1}\}.$$

Hence since  $E\{\xi^2\} < \infty$  and  $E\{\xi\} = 1$ , we have for  $|r| \leq 1, r \neq 1, r \sim 1$ ,

$$\begin{aligned} \frac{dP_1(r, 1)}{dr} [E\{(\xi-1)\xi\}(1-P_1(r, 1)) + o(1-P_1(r, 1))] = \\ \frac{1}{r}(1-E\{(\xi+1)\})(1-P_1(r, 1)) + \\ \frac{1}{2r} E\{(\xi+1)\xi\}(1-P_1(r, 1))^2 + o((1-P_1(r, 1))^2), \end{aligned}$$

so

$$\begin{aligned} \frac{dP_1(r, 1)}{dr} \left[ \alpha_1 + \frac{o(1-P_1(r, 1))}{1-P_1(r, 1)} \right] = \frac{1}{r} \frac{1}{1-P_1(r, 1)} - \frac{2}{r} + \\ \frac{\alpha_1 + 2}{2r} (1-P_1(r, 1)) + o(1-P_1(r, 1)). \end{aligned}$$

Inserting (3.5)i in the latter relation leads directly to

$$\frac{dP_1(r, 1)}{dr} \left[ 1 + \frac{o(\sqrt{1-r})}{\sqrt{1-r}} \right] = \frac{1}{2} \sqrt{2} \alpha_1^{-1/2} \frac{1}{\sqrt{1-r}} - \frac{2}{\alpha_1} + O(\sqrt{1-r}),$$

or

$$\frac{dP_1(r, 1)}{dr} = \frac{1}{2} \sqrt{2} \alpha_1^{-1/2} \frac{1}{\sqrt{1-r}} \left\{ 1 - \frac{o(\sqrt{1-r})}{\sqrt{1-r}} \right\} - \frac{2}{\alpha_1} + O(\sqrt{1-r}). \quad \square$$

LEMMA 3.3.

$$i. \quad \frac{dP_1(1, p_2)}{dp_2} \Big|_{p_2=1} := \lim_{p_2 \rightarrow 1} \frac{dP_1(1, p_2)}{dp_2} = \frac{1}{\alpha_1} \{ -\alpha_{12} \pm \sqrt{\alpha_{12}^2 - \alpha_1 \alpha_2} \}, \quad (3.6)$$

where the  $\pm$  sign corresponds to  $p_2 = e^{\pm i\phi}$  for  $\phi \downarrow 0$ ;

ii. for  $|p_2| = 1, p_2 \neq 1$  and  $p_2 \rightarrow 1$ ,

$$(1-p_2) \frac{d^2 P_1(1, p_2)}{dp_2^2} \rightarrow 0.$$

PROOF. With  $p_2 = e^{i\phi}, -\pi \leq \phi \leq \pi$ , it follows readily from lemma 3.1iii and

$$[1 - (1-p_2)][1 - (1-P_1(1, p_2))] - E\{[1 - (1-P_1(1, p_2))]^\xi [1 - (1-p_2)]^\eta\} = 0,$$

that for  $\phi \rightarrow \pm 0$ ,

$$\alpha_1(1 - P_1(1, p_2))^2 + 2\alpha_{12}(1 - p_2)(1 - P_1(1, p_2)) + \alpha_2(1 - p_2)^2 + o((1 - P_1(1, p_2))^2) + o((1 - p_2)^2) = 0.$$

Hence

$$\frac{1 - P_1(1, p_2)}{1 - p_2}$$

has a limit for  $\phi \downarrow 0$  as well as for  $\phi \uparrow 0$ . For  $|p_2| = 1, p_2 \neq 1$  we have from lemma 3.1,

$$\frac{dP_1(1, p_2)}{dp_2} = - \frac{P_1(1, p_2) E\{(1 - \eta)P_1^\xi(1, p_2)p_2^\eta\}}{p_2 E\{(1 - \xi)P_1^\xi(1, p_2)p_2^\eta\}}. \quad (3.7)$$

Obviously  $p_2 = 1$  is a zero of the numerator as well as at the denominator in (3.7), hence for  $p_2 \rightarrow 1$ ,

$$\frac{dP_1(1, p_2)}{dp_2} \Big|_{p_2=1} = - \frac{E\{(1 - \eta)\xi\} \frac{dP_1(1, p_2)}{dp_2} \Big|_{p_2=1} + E\{(1 - \eta)\eta\}}{E\{(1 - \xi)\xi\} \frac{dP_1(1, p_2)}{dp_2} \Big|_{p_2=1} + E\{(1 - \xi)\eta\}};$$

from which it follows

$$\alpha_1 \left[ \frac{dP_1(1, p_2)}{dp_2} \Big|_{p_2=1} \right]^2 + 2\alpha_{12} \frac{dP_1(1, p_2)}{dp_2} \Big|_{p_2=1} + \alpha_2 = 0, \quad (3.8)$$

and the latter relation leads directly to (3.6)i.

From (3.6)i it follows that  $p_2 = 1$  is a zero with multiplicity one of the numerator and denominator in (3.7). Since, cf. (3.6),

$$\begin{aligned} \frac{d}{dp_2} E\{(1 - \xi)P_1^\xi(1, p_2)p_2^\eta\} \Big|_{p_2=1} &= E\{(1 - \xi)\xi\} \frac{dP_1(1, p_2)}{dp_2} \Big|_{p_2=1} + E\{(1 - \xi)\eta\} \\ &= \mp \sqrt{\alpha_{12}^2 - \alpha_1\alpha_2} \neq 0, \end{aligned}$$

it is seen that (3.6)ii holds.  $\square$

LEMMA 3.4. For  $|r| \leq 1, r \sim 1$ ,

$$\frac{\partial P_1(r, p_2)}{\partial p_2} \Big|_{p_2=1} = - \frac{\alpha_{12}}{\alpha_1} + o(1). \quad (3.9)$$

PROOF. From : for  $|r| \leq 1, |p_2| = 1$ ,

$$p_2 P_1(r, p_2) = r E\{P_1^\xi(r, p_2)p_2^\eta\},$$

it follows readily, cf. lemma 3.liv that for  $|r| \leq 1, r \neq 1, |p_2| = 1$ ,

$$p_2 \frac{\partial P_1(r, p_2)}{\partial p_2} = -P_1(r, p_2) \frac{E\{(1-\eta)P_1^\xi(r, p_2)p_2^\eta\}}{E\{(1-\xi)P_1^\xi(r, p_2)p_2^\eta\}},$$

so for  $p_2 = 1$ ,

$$\left. \frac{\partial P_1(r, p_2)}{\partial p_2} \right|_{p_2=1} = -P_1(r, 1) \frac{E\{(1-\eta)P_1^\xi(r, 1)\}}{E\{(1-\xi)P_1^\xi(r, 1)\}}.$$

Hence for  $|r| \leq 1$ ,  $r \sim 1$ , by using (3.5)i,

$$\begin{aligned} \frac{\partial P_1(r, p_2)}{\partial p_2} &= -[1 - (1 - P_1(r, 1))] \frac{E\{(1-\eta)[1 - (1 - P_1(r, 1))]^\xi\}}{E\{(1-\xi)[1 - P_1(r, 1)]^\xi\}} \\ &= -[1 - (1 - P_1(r, 1))] \frac{E\{(1-\eta)\xi\}(1 - P_1(r, 1)) + o(1 - P_1(r, 1))}{E\{(1-\xi)\xi(1 - P_1(r, 1)) + o(1 - P_1(r, 1))\}} \\ &= -\frac{\alpha_{12}}{\alpha_1} + o(1). \quad \square \end{aligned}$$

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